Moment Generating Functions (MGFs)

Definition 1.1 The i^{th} moment of a random variable, X, is

$$\mathbf{E}_X(X^i) = \mu'_i \tag{1}$$

Definition 1.2 The *i*th central moment of a random variable, X, is

$$\mathbf{E}_X\left[(X-\mu)^i\right] = \mu_i \tag{2}$$

Note that $\mu = E_X(X^1)$ is the first moment of X, and that

$$\sigma^2 = \mathbf{E}\left[(X - \mathbf{E}_X(X))^2\right] \tag{3}$$

$$= E_X \left[(X - \mu)^2 \right] \tag{4}$$

is the second central moment of X.

Further, $\sigma^2 = E_X(X^2) - [E_X(X)]^2$ is expressible as a function of noncentral moments.

Definition 1.3 Suppose that X is a random variable with a given pmf, and that, for some real h > 0, $E_X(e^{sX})$ exists for every value of $s \in (-h, h)$. Then the function M defined by

$$M(s) = \mathcal{E}_X(e^{sX}) \tag{5}$$

is called the moment generating function of X.

Hence,

$$M(s) = \begin{cases} \sum_{x} e^{sx} p_X(x) & \text{discrete} \\ \int_{-\infty}^{\infty} e^{sx} f_X(x) dx \end{cases}$$
(6)

Why do we care? Recall that

$$e^x = 1 + x + \frac{x^2}{2!} + \frac{x^3}{3!} + \cdots$$
 (7)

Hence,

$$e^{sX} = \sum_{r=0}^{\infty} \frac{(sX)^r}{r!} \tag{8}$$

and thus,

$$M(s) = \mathcal{E}_X \left[\sum_{r=0}^{\infty} \frac{(sX)^r}{r!} \right]$$
(9)

$$= \sum_{r=0}^{\infty} \mathcal{E}_X(X^r) \frac{s^r}{r!} \tag{10}$$

We now note that the coefficient of $\frac{s^r}{r!}$ is $E_X(X^r)$. Now, we also note that, by Maclaurin series expansion,

$$M(s) = \sum_{r=0}^{\infty} \frac{d^r}{ds^r} M(s) \bigg|_{s=0} \frac{s^r}{r!}$$
(11)

$$= \sum_{r=0}^{\infty} M^{(r)}(0) \cdot \frac{s^r}{r!}$$
(12)

Comparing the coefficients of $\frac{s^r}{r!}$ we see that

$$E_X(X^r) = M^{(r)}(0) (13)$$

$$= \left. \frac{d^r}{ds^r} M(s) \right|_{s=0} \tag{14}$$

Hence, given an mgf we may generate moments until we are blue in the face.

Example 1.1 Consider $X \sim Bernoulli(p)$.

$$M(s) = e^{s(0)} p_X(0) + e^{s(1)} p_X(1)$$
(15)

$$= 1(1-p) + e^{s}(p) \tag{16}$$

$$= pe^{s} + 1 - p \quad -\infty < s < \infty \tag{17}$$

Thus,

$$\mathbf{E}_X(X) = \left. \frac{d}{ds} M(s) \right|_{s=0} \tag{18}$$

$$= p e^{s}|_{s=0} \tag{19}$$

$$= p \tag{20}$$

$$E_X(X^2) = \left. \frac{a^2}{ds^2} M(s) \right|_{s=0}$$
(21)

$$= p e^{s}|_{s=0} \tag{22}$$

$$= p \tag{23}$$

$$E_X(X^3) = \left. \frac{d^3}{ds^3} M(s) \right|_{s=0}$$
(24)

$$= p e^{s}|_{s=0} \tag{25}$$

$$= p \tag{26}$$

This simplifies the computation of

$$\operatorname{Var}(X) = p - p^2 \tag{27}$$

$$= p(1-p) \tag{28}$$

Example 1.2 Suppose $X \sim N(0, 1)$. Then

$$M(s) = \int_{-\infty}^{\infty} e^{sx} \frac{1}{\sqrt{2\pi}} e^{-x^2/2} dx$$
 (29)

$$= \frac{1}{\sqrt{2\pi}} \int_{-\infty}^{\infty} e^{-(x^2 - 2sx)/2} dx$$
 (30)

$$= \frac{1}{\sqrt{2\pi}} \int_{-\infty}^{\infty} e^{\left[-(x^2 - 2sx + s^2)/2\right] + s^2/2} dx \tag{31}$$

$$= e^{s^2/2} \frac{1}{\sqrt{2\pi}} \int_{-\infty}^{\infty} e^{-(x-s)^2/2} dx$$
 (32)

$$= e^{s^2/2} - \infty < s < \infty \tag{33}$$

Characteristics of MGFs 1.1

Theorem 1.1 Suppose a random variable, X, has the mgf M_X . Let Y = aX + b, for $a \in \Re$ and $b \in \Re$. Then the mgf of Y is

$$M_Y(s) = M_{aX+b}(s) \tag{34}$$

$$= e^{bs} M_X(as) \tag{35}$$

Proof:

$$M_{aX+b}(s) = \mathcal{E}_X\left(e^{(aX+b)s}\right) \tag{36}$$

$$= E_X \left(e^{bs} \cdot e^{asX} \right) \tag{37}$$

$$= e^{bs} \mathcal{E}_X \left(e^{asX} \right) \tag{38}$$

$$= e^{bs} M_X(as) \tag{39}$$

Example 1.3 Consider $Y \sim N(\mu, \sigma^2)$. Find the mgf of Y. Solution: Let $Z = \frac{Y-\mu}{\sigma}$. Then $Z \sim N(0, 1)$. Now, $Y = \mu + \sigma X$, and thus

$$M_Y(s) = M_{\mu+\sigma X}(s) \tag{40}$$

$$= e^{\mu s} M_X(\sigma s) \tag{41}$$

But, since $Z \sim N(0,1)$, $M_X(s) = e^{s^2/2}$ and

$$M_Y(s) = e^{\mu s} e^{(\sigma^2 s^2/2)}$$
(42)

$$= e^{\mu s + (\sigma^2 s^2/2)} \quad -\infty < s < \infty \tag{43}$$

Theorem 1.2 If X and Y are independent random variables with mgf's $M_X(s)$ and $M_Y(s)$ then for $a, b \in \Re$

$$M_{aX+bY}(s) = M_X(as) \cdot M_Y(bs) \tag{44}$$

Example 1.4 Recall that if $X \sim B(n, p)$ then $X = Y_1 + Y_2 + \cdots + Y_n$ where $Y_i \stackrel{iid}{\sim} Bernoulli(p)$. Since $M_{Y_i}(s) = pe^s + (1-p)$ for $i = 1, /2, \ldots, n$,

$$M_X(s) = M_{Y_1}(s)M_{Y_2}(s)\cdots M_{Y_n}(s)$$
(45)

 $= [pe^{s} + (1-p)]^{n}$ (46)

This points toward a general result which is of major importance.

Theorem 1.3 Two random variables share a common mgf if and only if they share a common distribution.

I.e. moment generating functions are unique.